

Introduction to targeted maximum likelihood estimation

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Notations

Plug-in estimation

One-step estimation

Targeted maximum likelihood estimation

Numerical application

Complete data structure and parameter of interest

$$\mathbb{D} = (X, Y(0), Y(1), A, Y) \sim \mathbb{P}$$

- ▶ X : vector of baseline covariates
- ▶ $Y(a)$, for $a \in \{0, 1\}$: potential outcome had the patient undergone treatment a (typically unobserved)
- ▶ A : binary treatment assignment
- ▶ $Y = AY(1) + (1 - A)Y(0)$: factual outcome (observed)

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Let us consider the average treatment effect (ATE) as target parameter

$$\mathbb{P} \mapsto E_{\mathbb{P}}[Y(1) - Y(0)]$$

defined in terms of potential outcomes

Observational data and identifiability

Observational data: n i.i.d. samples $\mathcal{D}_1, \dots, \mathcal{D}_n \sim P \in \mathcal{M}$

A generic data structure $\mathcal{D} = (X, A, Y)$:

Relevant features:

- ▶ $\mu_P(X, a) = E_P[Y|X, A = a]$
- ▶ $e_P(X, a) = P(A = a|X)$ (propensity score)

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$$\psi(P) \mapsto E_P [\mu_P(X, 1) - \mu_P(X, 0)]$$

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Our goal: build a good estimator of $\psi(P)$ using $\mathcal{D}_1, \dots, \mathcal{D}_n$

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Plug-in estimator

Plug-in estimators are **substitution estimators**

Learn the nuisance parameters: build an estimator μ_n of μ_P and define the plug-in estimator of $\psi(P)$ as

$$\psi_n = \frac{1}{n} \sum_{i=1}^n \mu_n(X_i, 1) - \mu_n(X_i, 0)$$

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Using a first-order expansion, the plug-in estimator can be decomposed as:

$$\psi_n = \psi(P) - E_P[\varphi_{\psi,n}(\mathcal{D})] + R(P, P_n),$$

where:

- ▶ $\varphi_{\psi,n}$ estimates the **efficient influence function (EIF)**
Intuition: A function describing how the estimator behaves under slight perturbations of the data-generating distribution

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- ▶ $R(P, P_n)$ is a second order remainder term, typically $o_P(1/\sqrt{n})$

Plug-in bias

Under the assumption that $\varphi_{\psi,n}(a)$ is a $L^2(P)$ -consistent estimator of $\varphi_{\psi}(P)$, and (b) lives in a P -Donsker class, it holds that

$$\begin{aligned}\psi_n - \psi(P) &= -E_P[\varphi_{\psi,n}(\mathcal{D})] + R(P, P_n) \\ &= -\frac{1}{n} \sum_{i=1}^n \varphi_{\psi,n}(\mathcal{D}_i) + \frac{1}{n} \sum_{i=1}^n \varphi_{\psi}(P)(\mathcal{D}_i) + o_P(1/\sqrt{n}) + R(P, P_n)\end{aligned}$$

Asymptotically linear estimator properties

Definition 1

An estimator ψ_n based on $\mathcal{D}_1, \dots, \mathcal{D}_n$ is **asymptotically linear** with influence function IF if it can be written as

$$\psi_n = \psi(P) + \frac{1}{n} \sum_{i=1}^n IF(\mathcal{D}_i) + o_P(1/\sqrt{n})$$

with $IF : \mathcal{D} \rightarrow \mathbb{R}$ such that $E_P[IF(\mathcal{D})] = 0$, $\text{Var}_P(IF(\mathcal{D})) < \infty$.
Asymptotically linear estimators are asymptotically normal:

$$\sqrt{n}(\psi_n - \psi(P)) \underset{n \rightarrow \infty}{\overset{P}{\rightsquigarrow}} \mathcal{N}(0, \text{Var}_P(IF(\mathcal{D})))$$

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$$\psi_n = \psi(P) - \frac{1}{n} \sum_{i=1}^n \varphi_{\psi, n}(\mathcal{D}_i) + \frac{1}{n} \sum_{i=1}^n \underbrace{\varphi_{\psi}(P)(\mathcal{D}_i)}_{IF(\mathcal{D}_i)} + o_P(1/\sqrt{n}) + R(P, P_n)$$

Plug-in estimator properties

- ▶ Plug-in estimators are **substitution estimators**
- ▶ The **random term**, is known and can drive the asymptotic behavior of ψ_n
- ▶ This motivates bias-corrected estimators such as the one-step estimator and TMLE

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From plug-in bias to one step correction

Intuition: Update ψ_n by directly correcting the first order **plug-in bias** in the **parameter space**

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 -\frac{1}{n} \sum_{i=1}^n \varphi_{\psi,n}(\mathcal{D}_i) &= -\frac{1}{n} \sum_{i=1}^n \left(\frac{A_i}{e_n(X_i, 1)} + \frac{1 - A_i}{e_n(X_i, 0)} \right) \cdot (Y_i - \mu_n(X_i, A_i)) \\
 &\quad + \mu_n(X_i, 1) - \mu_n(X_i, 0) - \psi_n \\
 &= -\frac{1}{n} \sum_{i=1}^n \left(\frac{A_i}{e_n(X_i, 1)} + \frac{1 - A_i}{e_n(X_i, 0)} \right) \cdot (Y_i - \mu_n(X_i, A_i))
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 \end{aligned}$$

The one-step estimator is defined as

$$\psi_n^{\text{os}} = \psi_n + \frac{1}{n} \sum_{i=1}^n \left(\frac{A_i}{e_n(X_i, 1)} + \frac{1 - A_i}{e_n(X_i, 0)} \right) \cdot (Y_i - \mu_n(X_i, A_i))$$

One-step estimator

Therefore,

$$\begin{aligned}\psi_n^{\text{OS}} - \psi(P) &= \psi_n - \psi(P) + \frac{1}{n} \sum_{i=1}^n \varphi_{\psi,n}(D_i) \\ &= -E_P[\varphi_{\psi,n}(\mathcal{D})] + R(P, P_n) + \frac{1}{n} \sum_{i=1}^n \varphi_{\psi,n}(D_i) \\ &= \frac{1}{n} \sum_{i=1}^n \varphi_{\psi}(P)(D_i) + o_P(1/\sqrt{n}) + R(P, P_n)\end{aligned}$$

One-step estimator properties

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By the central limit theorem,

$$\sqrt{n}(\psi_n^{\text{os}} - \psi(P)) \underset{n \rightarrow \infty}{\overset{P}{\rightsquigarrow}} \mathcal{N}(0, \text{Var}_P(\varphi_\psi(P)(\mathcal{D})))$$

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Double robust estimator:

$$R(P, P_n)^2 \leq \|\mu_n - \mu_P\|_P^2 \times \|(e_n - e_P)/\ell e_P\|_P^2,$$

if either $\mu_n = \mu_P$ or $e_P = e_n$, $R(P, P_n) = 0$ giving

$$\psi_n^{\text{OS}} = \psi(P) + \frac{1}{n} \sum_{i=1}^n \varphi_\psi(P)(\mathcal{D}_i) + o_P(1/\sqrt{n})$$

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$$\psi_n^{\text{OS}} = \psi(P) + \frac{1}{n} \sum_{i=1}^n \varphi_\psi(P)(\mathcal{D}_i) + o_P(1/\sqrt{n})$$



For binary outcomes, ψ_n^{OS} can step outside $[-1, 1]$
(due to propensity score weighting $\cdot/e_n(X, a)$ for $a \in \{0, 1\}$)

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Targeted learning motivation

Targeted maximum likelihood estimation (TMLE) performs the correction in the **nuisance model space**.

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How? Update μ_n into a targeted estimator μ_n^* , hence, building an update $\varphi_{\psi,n}^*$ of $\varphi_{\psi,n}$ (which also estimates $\varphi_{\psi}(P)$) such that

$$\frac{1}{n} \sum_{i=1}^n \varphi_{\psi,n}^*(\mathcal{D}_i) \approx 0$$

Targeted maximum likelihood estimation (TMLE)

Targeted maximum likelihood estimation (TMLE) proceeds in two steps:

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Targeting intuition

Can be viewed as regressing the residuals on the **clever covariate** (i.e. small fluctuation of μ_n along a direction given by the EIF)

$$Y - \mu_n(X, A) \approx \epsilon H(X, A)$$

where $H(X, A)$ is a function of e_n , defined from $\varphi_{\psi, n}(\mathcal{D}_i)$

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$$\underbrace{\left(\frac{A_i}{e_n(X_i, 1)} - \frac{1 - A_i}{e_n(X_i, 0)} \right)}_{=H(X, A)} (Y_i - \mu_n(X_i, A_i)) + \mu_n(X_i, 1) - \mu_n(X_i, 0) - \psi_n$$

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2. **Targeting step:** update μ_n into μ_n^* to mitigate the first order plug-in bias
 - a. Define a parametric submodel $\mu_{n,\epsilon}$
(i.e. express Y using μ_n , $H(X, A)$ and $\epsilon \in \mathbb{R}$)
 - b. Minimize a measure of discrepancy between Y and $\mu_{n,\epsilon}$
 $\hat{\epsilon} \in \operatorname{argmin}_{\epsilon \in \mathbb{R}} \ell_n(\epsilon)$ to define $\mu_n^* = \mu_{n,\hat{\epsilon}}$

Focus on step 2 (1/2)

2. Update μ_n into μ_n^* to correct plug-in bias (**continuous outcomes**):
 - a. Define a parametric submodel

$$\mu_{n,\epsilon}(X, A) = \mu_n(X, A) + \epsilon H(X, A)$$

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$$\hat{\epsilon} \in \operatorname{argmin}_{\epsilon \in \mathbb{R}} -\frac{1}{n} \sum_{i=1}^n [Y_i - \mu_{n,\epsilon}(X_i, A_i)]^2$$

which is such that

$$\frac{1}{n} \sum_{i=1}^n \varphi_{\psi^*,n}^*(\mathcal{D}_i) \approx 0$$

Proof

The derivative of $\mu_{n,\epsilon}(X, A) = \mu_n(X, A) + \epsilon H(X, A)$ w.r.t. ϵ :

$$\frac{d}{d\epsilon} \mu_{n,\epsilon}(X, A) = H(X, A)$$

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Therefore, $\hat{\epsilon} \in \operatorname{argmin}_{\epsilon \in \mathbb{R}} \ell_n(\epsilon)$ solves $\frac{d}{d\epsilon} \ell_n(\epsilon) = 0$:

$$\frac{d}{d\epsilon} \ell_n(\epsilon) = \frac{2}{n} \sum_{i=1}^n (Y_i - \mu_{n,\epsilon}(X_i, A_i)) H(X_i, A_i)$$

giving the desired result

Focus step 2 (2/2)

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$$\mu_{n,\epsilon}(X, A) = \text{expit}(\text{logit}(\mu_n(X, A)) + \epsilon H(X, A))$$

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$$\hat{\epsilon} \in \underset{\epsilon \in \mathbb{R}}{\text{argmin}} - \frac{1}{n} \sum_{i=1}^n [Y_i \log(\mu_{n,\epsilon}(X_i, A_i)) + (1 - Y_i) \log(1 - \mu_{n,\epsilon}(X_i, A_i))]$$

such that

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Therefore, $\hat{\epsilon} \in \text{argmin}_{\epsilon \in \mathbb{R}} \ell_n(\epsilon)$ solves $\frac{d}{d\epsilon} \ell_n(\epsilon) = 0$:

$$\begin{aligned} \frac{d}{d\epsilon} \ell_n(\epsilon) &= -\frac{1}{n} \sum_{i=1}^n Y_i \frac{\frac{d}{d\epsilon} \mu_{n,\epsilon}(X_i, A_i)}{\mu_{n,\epsilon}(X_i, A_i)} - (1 - Y_i) \frac{\frac{d}{d\epsilon} \mu_{n,\epsilon}(X_i, A_i)}{1 - \mu_{n,\epsilon}(X_i, A_i)} \\ &= -\frac{1}{n} \sum_{i=1}^n Y_i H(X_i, A_i) (1 - \mu_{n,\epsilon}(X_i, A_i)) - (1 - Y_i) H(X_i, A_i) \mu_{n,\epsilon}(X_i, A_i) \\ &= -\frac{1}{n} \sum_{i=1}^n (Y_i - \mu_{n,\epsilon}(X_i, A_i)) H(X_i, A_i) \end{aligned}$$

which gives the desired result

TMLE estimator

Define the TMLE estimator as

$$\psi_n^* = \frac{1}{n} \sum_{i=1}^n \mu_n^*(X_i, 1) - \mu_n^*(X_i, 0)$$

Continuous case:

$$\mu_n^*(X, 1) = \mu_n(X, 1) + \hat{\epsilon}H(X, 1)$$

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Therefore,

$$\psi_n^* - \psi(P) = -E_P[\varphi_{\psi^*, n}^*] + R(P, P_n^*)$$

$$= \underbrace{-\frac{1}{n} \sum_{i=1}^n \varphi_{\psi^*, n}^*(\mathcal{D}_i)}_{\approx 0} + \frac{1}{n} \sum_{i=1}^n \varphi_{\psi}(P)(\mathcal{D}_i) + o_P(1/\sqrt{n}) + R(P, P_n^*)$$

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Let's switch to R...

Thank you!

TMLE References: Books

1. Rubin, D., & van der Laan, M. (2006). Targeted Maximum Likelihood Learning. *The International Journal of Biostatistics*, 2(1), 1–38. <https://doi.org/10.2202/1557-4679.1043>
2. Laan, Mark J. van der, and Sherri Rose. 2011. *Targeted Learning*. Springer Series in Statistics. Springer, New York. <https://doi.org/10.1007/978-1-4419-9782-1>.
3. *Targeted Learning in Data Science*. Springer Series in Statistics. Springer, Cham. <https://doi.org/10.1007/978-3-319-65304-4>

TMLE References: Practical guides with R code

1. [A Ride in Targeted Learning Territory](#)
2. [An Illustrated Guide to TMLE](#)

References: R-packages

There are several available R-packages

- ▶ tmle

Gruber, S., & Laan, M. van der. (2012). tmle: An R Package for Targeted Maximum Likelihood Estimation. *Journal of Statistical Software*, 51(13), 1–35. <https://doi.org/10.18637/jss.v051.i13>

- ▶ tmle3